

# 17. R Analysis of Split-Plot Experiments

# Field Split-Plot Experiment

Field

Block 1	Genotype C				Genotype A				Genotype B			
	0	100	150	50	50	100	150	0	150	100	50	0
Block 2	Genotype B				Genotype A				Genotype C			
	150	100	50	0	0	50	150	100	100	50	150	0
Block 3	Genotype A				Genotype B				Genotype C			
	100	50	0	150	0	100	150	50	50	100	150	0
Block 4	Genotype B				Genotype C				Genotype A			
	0	50	100	150	150	100	50	0	50	150	100	0

# Install and Load Packages for Linear Mixed-Model Analysis in R

```
> install.packages("lme4")
```

```
> install.packages("lmerTest")
```

```
> library(lme4)
```

```
> library(lmerTest)
```

## Read Data

```
> fd = read.delim(  
  "https://dnett.github.io/S510/FieldSplitPlotData.txt")
```

```
> head(fd)
```

	block	geno	fert	y
1	1	1	0	148.7
2	1	1	50	150.4
3	1	1	100	166.7
4	1	1	150	156.5
5	1	2	0	162.5
6	1	2	50	168.6

```
> #Define factors and shorten names.
```

```
> y = fd$y
```

```
> b = factor(fd$block)
```

```
> g = factor(fd$geno)
```

```
> f = factor(fd$fert/50+1)
```

## Fit the Model

```
> o = lmer(y ~ g + f + g:f + (1 | b) + (1 | b:g))
```

$$y_{ijk} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \mathbf{b}_k + w_{ik} + e_{ijk}$$

$$\mathbf{b}_k \stackrel{iid}{\sim} N(\mathbf{0}, \sigma_b^2)$$

$$w_{ik} \stackrel{iid}{\sim} N(0, \sigma_w^2)$$

$$e_{ijk} \stackrel{iid}{\sim} N(0, \sigma_e^2)$$

```
> summary(o)
```

```
Linear mixed model fit by REML.
```

```
t-tests use Satterthwaite's method ['lmerModLmerTest']
```

```
Formula: y ~ g + f + g:f + (1 | b) + (1 | b:g)
```

```
REML criterion at convergence: 275.1
```

```
Scaled residuals:
```

Min	1Q	Median	3Q	Max
-1.43560	-0.60942	-0.07892	0.53263	2.09822

Random effects:

Groups	Name	Variance	Std.Dev.
b:g	(Intercept)	67.29	8.203
b	(Intercept)	122.89	11.086
Residual		39.71	6.301

Number of obs: 48, groups: b:g, 12; b, 4



## Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t )	
(Intercept)	126.025	7.581	6.740	16.624	1.01e-06	***
g2	22.500	7.314	11.154	3.076	0.010384	*
g3	-12.275	7.314	11.154	-1.678	0.121064	
f2	13.750	4.456	27.000	3.086	0.004649	**
f3	22.425	4.456	27.000	5.033	2.79e-05	***
f4	19.025	4.456	27.000	4.270	0.000216	***
g2:f2	-2.825	6.301	27.000	-0.448	0.657499	
g3:f2	17.875	6.301	27.000	2.837	0.008540	**
g2:f3	3.750	6.301	27.000	0.595	0.556721	
g3:f3	24.525	6.301	27.000	3.892	0.000588	***
g2:f4	8.600	6.301	27.000	1.365	0.183583	
g3:f4	35.800	6.301	27.000	5.681	4.92e-06	***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 '1'

Correlation of Fixed Effects:

	(Intr)	g2	g3	f2	f3	f4	g2:f2	g3:f2	g2:f3	g3:f3	g2:f4
g2	-0.482										
g3	-0.482	0.500									
f2	-0.294	0.305	0.305								
f3	-0.294	0.305	0.305	0.500							
f4	-0.294	0.305	0.305	0.500	0.500						
g2:f2	0.208	-0.431	-0.215	-0.707	-0.354	-0.354					
g3:f2	0.208	-0.215	-0.431	-0.707	-0.354	-0.354	0.500				
g2:f3	0.208	-0.431	-0.215	-0.354	-0.707	-0.354	0.500	0.250			
g3:f3	0.208	-0.215	-0.431	-0.354	-0.707	-0.354	0.250	0.500	0.500		
g2:f4	0.208	-0.431	-0.215	-0.354	-0.354	-0.707	0.500	0.250	0.500	0.250	
g3:f4	0.208	-0.215	-0.431	-0.354	-0.354	-0.707	0.250	0.500	0.250	0.500	0.500

```
> anova(o)
```

```
Type III Analysis of Variance Table with Satterthwaite's method
```

	Sum Sq	Mean Sq	NumDF	DenDF	F value	Pr(>F)	
g	673.3	336.65	2	6.0013	8.4784	0.0178463	*
f	8737.7	2912.57	3	26.9999	73.3529	4.234e-13	***
g:f	1557.3	259.56	6	26.9999	6.5369	0.0002381	***

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
> ls_means(o)
```

```
Least Squares Means table:
```

	Estimate	Std. Error	df	t value	lower	upper	Pr(> t )	
g1	139.825	7.073	5.1	19.769	121.779	157.870	4.870e-06	***
g2	164.706	7.073	5.1	23.287	146.661	182.752	2.120e-06	***
g3	147.100	7.073	5.1	20.797	129.054	165.145	3.765e-06	***
f1	129.433	6.296	3.4	20.558	110.696	148.171	0.000107	***
f2	148.200	6.296	3.4	23.539	129.463	166.937	6.758e-05	***
f3	161.283	6.296	3.4	25.617	142.546	180.020	5.071e-05	***
f4	163.258	6.296	3.4	25.930	144.521	181.995	4.865e-05	***

	Estimate	Std. Error	df	t value	lower	upper	Pr(> t )	
g1:f1	126.025	7.581	6.7	16.624	107.958	144.092	1.012e-06	***
g2:f1	148.525	7.581	6.7	19.592	130.458	166.592	3.411e-07	***
g3:f1	113.750	7.581	6.7	15.005	95.683	131.817	1.986e-06	***
g1:f2	139.775	7.581	6.7	18.437	121.708	157.842	5.102e-07	***
g2:f2	159.450	7.581	6.7	21.033	141.383	177.517	2.129e-07	***
g3:f2	145.375	7.581	6.7	19.176	127.308	163.442	3.932e-07	***
g1:f3	148.450	7.581	6.7	19.582	130.383	166.517	3.423e-07	***
g2:f3	174.700	7.581	6.7	23.044	156.633	192.767	1.159e-07	***
g3:f3	160.700	7.581	6.7	21.198	142.633	178.767	2.021e-07	***
g1:f4	145.050	7.581	6.7	19.133	126.983	163.117	3.991e-07	***
g2:f4	176.150	7.581	6.7	23.236	158.083	194.217	1.097e-07	***
g3:f4	168.575	7.581	6.7	22.236	150.508	186.642	1.470e-07	***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 '1'

Confidence level: 95%

Degrees of freedom method: Satterthwaite

```
> betahat=fixef(o)
```

```
> betahat
```

(Intercept)		g2		g3	
126.025		22.500		-12.275	
	f2		f3		f4
13.750		22.425		19.025	
g2:f2	g3:f2	g2:f3	g3:f3	g2:f4	g3:f4
-2.825	17.875	3.750	24.525	8.600	35.800

```

> #The following table shows the cell means
> #in terms of the R parameterization.
> #####
> #
> #   f1=0units  f2=50units      f3=100units      f4=150units
> #
> #g1  mu          mu    +f2          mu    +f3          mu    +f4
> #
> #g2  mu+g2       mu+g2+f2+g2f2      mu+g2+f3+g2f3      mu+g2+f4+g2f4
> #
> #g3  mu+g3       mu+g3+f2+g3f2      mu+g3+f3+g3f3      mu+g3+f4+g3f4
> #
> #####

```

```
> #Coefficients for geno 1 marginal mean
>
> C1 = matrix(c(1,
+             0, 0,
+             1/4, 1/4, 1/4,
+             0, 0, 0, 0, 0, 0), nrow=1)
>
```



```
> #Coefficients for geno 1 - geno 2 marginal mean
>
> C2 = matrix(c(0,
+             -1, 0,
+             0, 0, 0,
+             -1/4, 0, -1/4, 0, -1/4, 0), nrow=1)
>
>
```

```
> #Coefficients for geno 1 - geno 2 with no fertilizer
>
> C3 = matrix(c(0,
+             -1, 0,
+             0, 0, 0,
+             0, 0, 0, 0, 0, 0), nrow=1)
>
```

```
> C = rbind(C1, C2, C3)
> contest(o, L = C, joint = F, confint = T)
```

	Estimate	Std. Error	df	t value	lower	upper	Pr(> t )
1	139.82500	7.072992	5.129	19.768862	121.77948	157.870523	4.870369e-06
2	-24.88125	6.213453	6.001	-4.004416	-40.08425	-9.678254	7.080208e-03
3	-22.50000	7.314154	11.154	-3.076227	-38.57124	-6.428757	1.038405e-02

```
> #The degrees of freedom, sums of squares,  
> #and mean squares from a sequential ANOVA  
> #table could be used to estimate variance  
> #components and compute test statistics.  
>  
> a = anova(lm(y ~ b + g + b:g + f + g:f))
```

```
> a
```

```
Analysis of Variance Table
```

```
Response: y
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
b	3	5349.5	1783.16	44.9089	1.252e-10	***
g	2	5237.2	2618.62	65.9500	4.057e-11	***
f	3	8737.7	2912.57	73.3531	4.233e-13	***
b:g	6	1853.4	308.90	7.7796	6.355e-05	***
g:f	6	1557.3	259.56	6.5370	0.0002381	***
Residuals	27	1072.1	39.71			

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
> #For example, based on expected mean squares
> #presented in slide set 15, an unbiased
> #estimator of variance for the whole-plot
> #random effects is (MSbg - MSe) / 4.
>
> MSbg = a[4,3]
> MSbg
[1] 308.8985
> MSe = a[6,3]
> MSe
[1] 39.70613
> (MSbg - MSe) / 4
[1] 67.2981
```

```
> #The F tests and p-values in the
> #ANOVA table "a" all use MSe as the
> #denominator because the lm function
> #implies a Gauss-Markov model rather
> #than a linear mixed-effects model.
> #This is fine for f and g:f, but not
> #for the other lines of the ANOVA
> #table.
```

```
> #The correct F statistic for testing
> #for genotype main effects is
> #MSg/MSbg
>
> MSg = a[2,3]
> MSg
[1] 2618.619
>
> MSbg = a[4,3]
> MSbg
[1] 308.8985
>
> MSg / MSbg
[1] 8.47728
```



```
> #When the experiment is balanced, whole-plot-factor analysis
> #can be accomplished by computing the average for each
> #whole-plot experimental unit and then analyzing those
> #averages.
>
> d = aggregate(y, by = list(b, g), FUN = mean)
> names(d) = c("block", "geno", "wpaverage")
```

```
> d
```

```
  block geno wpaverage
1      1     1  155.575
2      2     1  134.075
3      3     1  135.100
4      4     1  134.550
5      1     2  173.100
6      2     2  162.925
7      3     2  154.600
8      4     2  168.200
9      1     3  177.250
10     2     3  142.800
11     3     3  137.325
12     4     3  131.025
```

```
> anova(lm(wpaverage ~ block + geno, data = d))
```

```
Analysis of Variance Table
```

```
Response: wpaverage
```

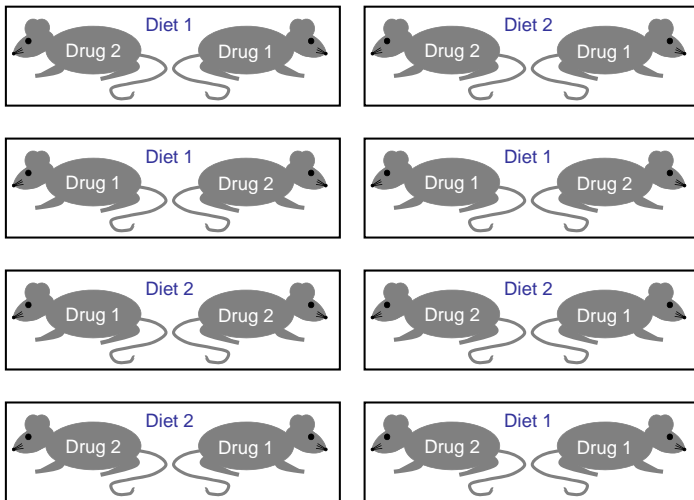
	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
block	3	1337.37	445.79	5.7726	0.03346	*
geno	2	1309.31	654.65	8.4773	0.01786	*
Residuals	6	463.35	77.22			

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
> #Now analyze the diet drug split-plot experiment.  
> #The main difference here is that the whole-plot  
> #part of the experiment is a CRD rather than an  
> #RCBD.
```

# Diet and Drug Split-Plot Experiment



```
> ddd = read.delim(
  "https://dnett.github.io/S510/DietDrugSplitPlotData.txt")
```

```
> head(ddd)
```

	litter	diet	drug	y
1	1	1	2	18.8
2	1	1	1	10.3
3	2	2	2	14.5
4	2	2	1	18.5
5	3	1	1	16.8
6	3	1	2	27.2

```
> ddd$litter = factor(ddd$litter)
> ddd$diet = factor(ddd$diet)
> ddd$drug = factor(ddd$drug)
>
> o = lmer(y ~ diet + drug + diet:drug + (1 | litter), data=ddd)
```

$$y_{ijk} = \mu + \alpha_i + \beta_j + \gamma_{ij} + \ell_{ik} + e_{ijk}$$

$$\ell_{ik} \stackrel{iid}{\sim} N(0, \sigma_\ell^2)$$

$$e_{ijk} \stackrel{iid}{\sim} N(0, \sigma_e^2)$$

```
> summary(o)
```

```
Linear mixed model fit by REML.
```

```
t-tests use Satterthwaite's method ['lmerModLmerTest']
```

```
Formula: y ~ diet + drug + diet:drug + (1 | litter)
```

```
Data: ddd
```

```
REML criterion at convergence: 57.4
```

```
Scaled residuals:
```

Min	1Q	Median	3Q	Max
-0.95525	-0.51772	0.05178	0.51149	1.02906



Random effects:

Groups	Name	Variance	Std.Dev.
litter	(Intercept)	9.7871	3.128
Residual		0.9408	0.970

Number of obs: 16, groups: litter, 8

Fixed effects:

	Estimate	Std. Error	df	t value	Pr(> t )	
(Intercept)	13.5750	1.6377	6.5492	8.289	0.000105	***
diet2	2.9250	2.3160	6.5492	1.263	0.249713	
drug2	9.7500	0.6859	6.0000	14.216	7.57e-06	***
diet2:drug2	-11.8500	0.9700	6.0000	-12.217	1.83e-05	***

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

Correlation of Fixed Effects:

	(Intr)	diet2	drug2	
diet2		-0.707		
drug2		-0.209	0.148	
diet2:drug2		0.148	-0.209	-0.707

```
> anova(o)
```

```
Type III Analysis of Variance Table with Satterthwaite's method
```

	Sum Sq	Mean Sq	NumDF	DenDF	F value	Pr(>F)	
diet	1.651	1.651	1	6	1.7548	0.2334898	
drug	58.522	58.522	1	6	62.2028	0.0002202	***
diet:drug	140.422	140.422	1	6	149.2533	1.831e-05	***

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
> ls_means(o)
```

```
Least Squares Means table:
```

	Estimate	Std. Error	df	t value	lower	upper	Pr(> t )	
diet1	18.4500	1.6014	6.0	11.5214	14.5316	22.3684	2.569e-05	***
diet2	15.4500	1.6014	6.0	9.6480	11.5316	19.3684	7.102e-05	***
drug1	15.0375	1.1580	6.5	12.9856	12.2606	17.8144	6.455e-06	***
drug2	18.8625	1.1580	6.5	16.2887	16.0856	21.6394	1.522e-06	***
diet1:drug1	13.5750	1.6377	6.5	8.2892	9.6478	17.5022	0.0001047	***
diet2:drug1	16.5000	1.6377	6.5	10.0753	12.5728	20.4272	3.169e-05	***
diet1:drug2	23.3250	1.6377	6.5	14.2428	19.3978	27.2522	3.589e-06	***
diet2:drug2	14.4000	1.6377	6.5	8.7930	10.4728	18.3272	7.320e-05	***

```
---
```

```
Signif. codes:  0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1
```

```
Confidence level: 95%
```

```
Degrees of freedom method: Satterthwaite
```

```

> #The following table shows the cell means
> #in terms of the R parameterization.
> #####
> #
> #           drug 1           drug 2
> #
> #diet 1      mu                mu + drug2
> #
> #diet 2      mu + diet2        mu + diet2 + drug2 + diet2:drug2
> #
> #####

```

```
> #Coefficients for diet 1 - diet 2 marginal mean
>
> C1 = matrix(c(0, -1, 0, -1/2), nrow = 1)
>
```

```
> #Coefficients for drug 1 - drug 2 marginal mean
>
> C2 = matrix(c(0, 0, -1, -1/2), nrow = 1)
>
```



```
> #Coefficients for diet 1 - diet 2 for drug 2
>
> C3 = matrix(c(0, -1, 0, -1), nrow=1)
>
```

```
> #Coefficients for drug 1 - drug 2 for diet 2
>
> C4 = matrix(c(0, 0, -1, -1), nrow=1)
>
```

```

> C = rbind(C1, C2, C3, C4)
> contest(o, L = C, joint = F, confint = T)

```

	Estimate	Std. Error	df	t value	lower	upper	Pr(> t )
1	3.000	2.2646744	6.000000	1.324694	-2.5414586	8.541459	0.23349
2	-3.825	0.4849828	6.000000	-7.886877	-5.0117102	-2.638290	0.00022
3	8.925	2.3160221	6.549174	3.853590	3.3711210	14.478879	0.00712
4	2.100	0.6858693	6.000000	3.061808	0.4217383	3.778262	0.02217

> #The F statistic for diet main effects  
> #is also given by the ratio  
> #MSdiet / MSlitter from the appropriate  
> #sequential ANOVA table. This can be  
> #shown formally, but MSlitter makes  
> #sense as the denominator mean square  
> #because litters are the experimental  
> #units for the factor diet.

```
> a = anova(
+   lm(y ~ diet + litter + drug + diet:drug,
+     data=ddd))
```

```
> a
```

### Analysis of Variance Table

Response: y

	Df	Sum Sq	Mean Sq	F value	Pr(>F)	
diet	1	36.000	36.000	38.264	0.0008214	***
litter	6	123.090	20.515	21.805	0.0007887	***
drug	1	58.523	58.523	62.203	0.0002202	***
diet:drug	1	140.423	140.423	149.253	1.831e-05	***
Residuals	6	5.645	0.941			

---

Signif. codes: 0 '\*\*\*' 0.001 '\*\*' 0.01 '\*' 0.05 '.' 0.1 ' ' 1

```
> MSdiet = a[1,3]
> dfdiet = a[1,1]
> MSlitter = a[2,3]
> dflitter = a[2,1]
> F = MSdiet / MSlitter
> p = 1 - pf(F, dfdiet, dflitter)
> F
[1] 1.754814
> p
[1] 0.2334898
```

```
> #When the experiment is balanced, whole-plot-factor analysis  
> #can be accomplished by computing the average for each  
> #whole-plot experimental unit and then analyzing those  
> #averages.
```

```
> d = aggregate(ddd$y, by = list(ddd$diet, ddd$litter),  
                FUN = mean)  
  
> names(d) = c("diet", "litter", "wpaverage")
```



```
> d
  diet litter wpaverage
1    1     1    14.55
2    2     2    16.50
3    1     3    22.00
4    1     4    16.00
5    2     5    11.60
6    2     6    17.00
7    2     7    16.70
8    1     8    21.25
```

```
> anova(lm(wpaverage ~ diet, data = d))
```

```
Analysis of Variance Table
```

```
Response: wpaverage
```

	Df	Sum Sq	Mean Sq	F value	Pr(>F)
diet	1	18.000	18.000	1.7548	0.2335
Residuals	6	61.545	10.258		