STAT 511 Exam 1 Spring 2010

- 1. Suppose y is an $n \times 1$ response vector and X is an $n \times p$ design matrix.
 - (a) State the Gauss-Markov linear model.
 - (b) Provide a matrix formula for the best linear unbiased estimator (BLUE) of E(y) in terms of X and y.
 - (c) State the normal equations.
 - (d) Suppose b is any solution to the normal equations. Is it necessarily true that the best linear unbiased estimator in part (b) is equal to Xb? Prove that your answer is correct.
- 2. Prove or disprove the following statement: Every symmetric and idempotent matrix is an orthogonal projection matrix.
- 3. Consider an experiment designed to study the effect of two dietary factors, protein source and protein amount, on weight gain in pigs. A total of 12 pigs were randomly assigned to treatment with one of six combinations of protein source (1 or 2) and protein amount (1, 2, or 3 units). A completely randomized design was used with two individually penned pigs per treatment group. Let y_{ijk} denote the amount of weight gained during the study period by the k^{th} pig fed j units of protein from source i (i = 1, 2; j = 1, 2, 3; k = 1, 2). Consider the model

$$y_{ijk} = \mu + s_i + \beta x_j + \epsilon_{ijk}, \ (i = 1, 2; j = 1, 2, 3; k = 1, 2)$$

where μ , s_1 , s_2 , and β are unknown real-valued parameters, $x_j = j - 2$, the ϵ_{ijk} 's are *iid* $N(0, \sigma^2)$, and σ^2 is an unknown parameter in \mathbb{R}^+ . Suppose

$$y = [y_{111}, y_{112}, y_{121}, y_{122}, \dots, y_{231}, y_{232}]',$$

$$\epsilon = [\epsilon_{111}, \epsilon_{112}, \epsilon_{121}, \epsilon_{122}, \dots, \epsilon_{231}, \epsilon_{232}]',$$

and $\beta = [\mu, s_1, s_2, \beta]'.$

- (a) Provide the appropriate design matrix X so that the model may be written as $y = X\beta + \epsilon$.
- (b) For each of the quantities below, state whether the quantity is estimable and prove that your answer is correct.
 - i. $\mu + s_1$ ii. $\mu + s_1 + 10\beta$ iii. $s_1 - s_2$ iv. μ
- (c) Write down a full-column-rank matrix that has the same column space as X in part (a).
- (d) Use your answer to part (c) to find a simplified expression for the BLUE of $E(y_{111})$ in terms of the y_{ijk} values.
- (e) Provide the least squares estimate of each estimable quantity in part (b).

4. Once again consider the experiment described in problem 3. Consider the model

$$y_{ijk} = \mu_{ij} + \epsilon_{ijk}$$
 $(i = 1, 2; j = 1, 2, 3; k = 1, 2)$

where $\mu_{11}, \ldots, \mu_{23}$ are unconstrained, unknown, real-valued cell mean parameters; the ϵ_{ijk} 's are *iid* $N(0, \sigma^2)$; and σ^2 is an unknown parameter in \mathbb{R}^+ . Use the R code and output provided on the last page of this exam to complete the following parts.

- (a) Is there evidence of a difference among the six treatment means? Provide a test statistic, its degrees of freedom, a *p*-value, and a conclusion.
- (b) Provide the BLUE of μ_{23} .
- (c) If possible, provide the standard error for the estimate in part (b). If it is not possible to determine the standard error using the information provided, explain why.
- (d) Provide a matrix C so that $C \approx \approx \text{coef}(o)$ is an estimate of the main effect of protein source.
- (e) If the model in problem 3 were fit to these data, what would the estimate of the error variance σ^2 be?
- (f) Suppose the researchers would like to know if the model specified in the statement of problem 3 fits these data adequately relative to the cell means model specified in the statement of problem 4. Compute a test statistic that can be used to answer this question and state the degrees of freedom associated with this test statistic.
- 5. The mean and variance of a χ_m^2 random variable are *m* and 2*m*, respectively. Suppose $w \sim \chi_m^2(\delta^2)$. Use the given information about central chi-square random variables to help you derive the following:
 - (a) E(w),
 - (b) Var(w). (To derive the variance, it might help you to know that $Cov(z, z^2) = 0$ if $z \sim N(0, 1)$.)

```
> y=#DATA NOT SHOWN#
> s=factor(rep(1:2,each=6))
> s
[1] 1 1 1 1 1 1 2 2 2 2 2 2 2
Levels: 1 2
> x = rep(rep(c(-1, 0, 1), each=2), 2)
> x
> o=lm(y \sim s + x + I(x^2) + s : x + s : I(x^2))
> anova(o)
Analysis of Variance Table
Response: y
         Df Sum Sq Mean Sq F value Pr(>F)
          1 90.75 90.75 27.9231 0.001858 **
S
          1 528.12 528.12 162.5000 1.430e-05 ***
х
I(x^2)
         1 0.04
                    0.04 0.0128 0.913544
         1 28.13 28.13 8.6538 0.025889 *
s:x
s:I(x^2) 1 0.37 0.37 0.1154 0.745670
Residuals 6 19.50
                    3.25
___
Signif. codes: 0 `***' 0.001 `**' 0.01 `*' 0.05 `.' 0.1 ` ' 1
> summary(o)
Call:
lm(formula = y \sim s + x + I(x^2) + s:x + s:I(x^2))
Residuals:
      Min
                  10
                        Median
                                      30
                                                Max
-2.000e+00 -1.125e+00 -1.712e-16 1.125e+00 2.000e+00
Coefficients:
           Estimate Std. Error t value Pr(>|t|)
(Intercept) 17.5000 1.2748 13.728 9.29e-06 ***
s2
             5.0000
                      1.8028 2.774 0.032273 *
             6.2500
                      0.9014 6.934 0.000446 ***
х
I(x^2)
           -0.2500
                      1.5612 -0.160 0.878035
                      1.2748 2.942 0.025889 *
             3.7500
s2:x
s2:I(x^2)
            0.7500
                      2.2079 0.340 0.745670
___
Signif. codes: 0 `***' 0.001 `**' 0.01 `*' 0.05 `.' 0.1 ` ' 1
Residual standard error: 1.803 on 6 degrees of freedom
Multiple R-squared: 0.9708, Adjusted R-squared: 0.9464
F-statistic: 39.84 on 5 and 6 DF, p-value: 0.0001587
```